

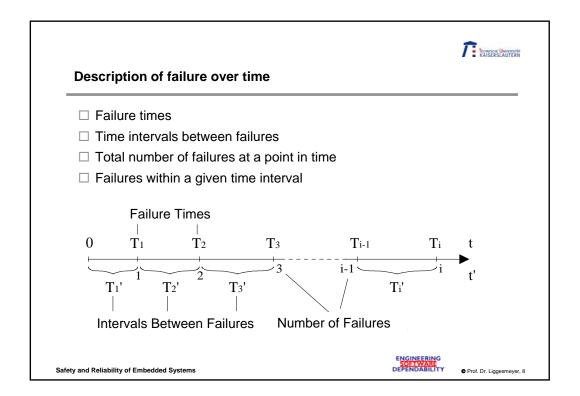


Tool Assisted Reliability Modeling

- ☐ Use of models
 - Which models do exist?
 - How can I find out, which model fits my purposes best?
 - How can I define the model parameters in order to get dependable reliability predictions?

Safety and Reliability of Embedded Systems







Modeling of Reliability

☐ Lifetime T

- Large number of similar systems under consideration
- Simultaneous start of the systems at time t = 0
- Observed time of the first failure of each system is the so-called lifetime T of this system
- Plot of the fraction of failed systems over t is the so-called empirical distribution function of the lifetime (or empirical life distribution)

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Modeling of Reliability

- ☐ If the number of systems becomes larger (approximates infinity), the empirical life distribution approximates the life distribution F(t)
 - Here, lifetime T is a random variable and F(t) is the probability that an arbitrary system is not operational at t F(t) = P{T ≤ t}
 - F(t) is the probability that lifetime T is less or equal to t, meaning that a system has already failed by t.
 - We use the following assumptions:

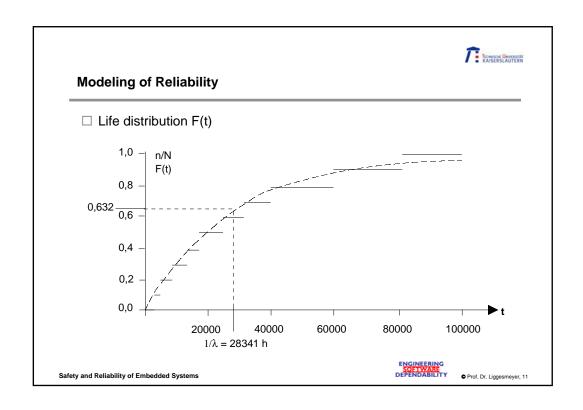
F(t = 0) = 0, i.e. a new system is intact, and $\lim_{t \to \infty} F(t) = 1$, i.e. every system fails sometimes

 $\hfill\Box$ Failure Times of 10 Systems

T _i (h)	2810	5411	8701	13130	17327	24899	31230	40006	59880	80017
n/ _N	0,1	0,2	0,3	0,4	0,5	0,6	0,7	0,8	0,9	1,0
F(t)	0,094	0,174	0,264	0,371	0,457	0,585	0,668	0,765	0,879	0,941

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Modeling of Reliability

- ☐ Reliability function R(t)
 - F(t) gives the probability that at time t at least one failure has occurred; thus R(t) = 1 - F(t) is the probability that at time t no failure has occurred yet
- ☐ Probability density f(t)
 - The probability density f(t) describes the modification of the probability that a system fails over time

$$f(t) = \frac{d F(t)}{dt}$$

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Modeling of Reliability

☐ MTBF, MTTF

- A relevant measure for reliability is the Mean Time To Failure (MTTF) or Mean Time Between Failure (MTBF)
- The MTTF resp. MTBF defines the mean value of the lifetime resp. the mean value for the time interval between two successive failures
- It is determined by calculating the following integral:

$$\overline{T} = E(T) = \int_{0}^{\infty} t f(t) dt$$

□ Failure rate

The failure rate is the relative boundary value of failed entities at time t in a time interval that approximates zero, referring to the entities still functional at the beginning of the time interval

$$\lambda(t) = \frac{f(t)}{R(t)} = \ \frac{dF(t) \ / \ dt}{R(t)} = \frac{- \ dR(t) \ / \ dt}{R(t)}$$

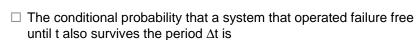
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Modeling of Reliability



$$\frac{\mathsf{R}(\mathsf{t} + \Delta \mathsf{t})}{\mathsf{R}(\mathsf{t})}$$

 \square Thus, the probability that the product fails within Δt is

$$1 - \frac{R(t + \Delta t)}{R(t)} = 1 - \frac{1 - F(t + \Delta t)}{1 - F(t)} = \frac{1 - F(t) - (1 - F(t + \Delta t))}{1 - F(t)} = \frac{F(t + \Delta t) - F(t)}{1 - F(t)}$$

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Modeling of Reliability

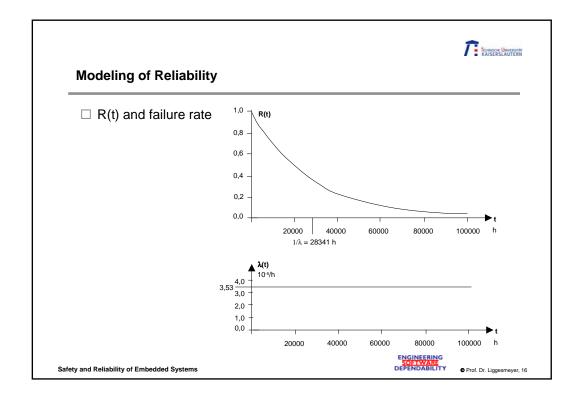
 \Box As the given probability for short time intervals Δt is proportional to $\Delta t,$ we divide the term by Δt and determine the boundary value when Δt approximates 0

$$\lim_{\Delta t \, \rightarrow \, 0} \ \frac{1}{\Delta t} \ \frac{F(t+\Delta t) - F(t)}{1 - F(t)} = \frac{1}{R(t)} \ \lim_{\Delta t \, \rightarrow \, 0} \ \frac{F(t+\Delta t) - F(t)}{\Delta t} = \frac{f(t)}{R(t)} = \lambda(t)$$

 \square Thus the probability that a system, that is operational at time t fails within the (short) time interval Δt , is approximately Δt $\lambda(t)$

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Example for the Distribution Function

□ Assumption: For the given data (table p. 10) lifetime is exponentially distributed: $F(t) = 1 - e^{-\lambda t}$

 \Box The parameter λ (failure rate) has to be determined based on failure observations in order to achieve an optimal adjustment of the function, according to a predetermined criterion. The Maximum-Likelihood-Method provides the following parameter λ for the exponential distribution:

$$\lambda = \frac{N}{\sum_{i=1}^{N} T_{i}} = 0,0000353 / h$$

□ Reliability: $R(t) = 1 - F(t) = e^{-\lambda t}$

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Example for the Distribution Function

 \Box The failure rate λ is constant over time

$$\lambda(t) = \frac{f(t)}{R(t)} = \frac{\mathrm{d}F(t) \, / \, \mathrm{d}t}{R(t)} = \frac{\mathrm{-\,d}R(t) \, / \, \mathrm{d}t}{R(t)} = \frac{\lambda \, e^{-\lambda t}}{e^{-\lambda t}} = \lambda$$

☐ A constant failure rate causes an exponential distribution of the lifetime

□ Determination of the MTTF

$$\overline{T} = E(T) = \int_{0}^{\infty} t f(t) dt = \int_{0}^{\infty} t \lambda e^{-\lambda t} dt = \lambda \int_{0}^{\infty} t e^{-\lambda t} dt = \lambda \left(\frac{\lambda e^{-\lambda t}}{\lambda^{2}} (-\lambda t - 1) \right) = \frac{1}{\lambda}$$

☐ If lifetime is exponentially distributed, the MTTF is the reciprocal of the failure rate and thus constant

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The Exponential Distribution

- □ Life distribution: $F(t) = 1 e^{-\lambda t}$
- □ Density function: $f(t) = \lambda e^{-\lambda t}$
- \Box Reliability function: R(t) = 1 F(t) = $e^{-\lambda t}$
- \Box Failure rate: $\lambda(t) = \lambda$
- \Box MTTF: $\overline{T} = \frac{1}{\lambda}$

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The Weibull Distribution

- $\hfill \Box$ Life Distribution : F(t) = 1 $e^{\text{-}(\lambda t)\beta}$; $\lambda,\,\beta>0$
- \square or:
- $F(t) = 1 e^{-\frac{1}{\alpha}t^{\beta}}; \alpha, \beta > 0, d. h. \frac{1}{\alpha} = \lambda^{\beta}$
- ☐ Density:

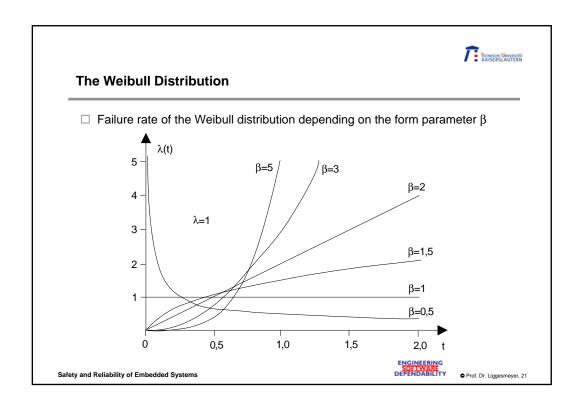
$$f(t) = \frac{dF(t)}{dt} = \lambda \beta (\lambda t)^{\beta-1} e^{-(\lambda t)^{\beta}}$$

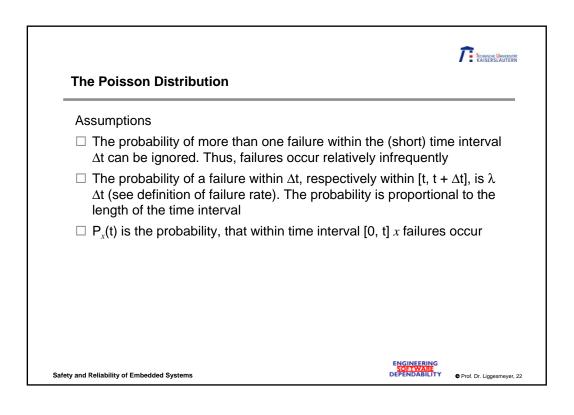
- \Box Reliability: R(t) = $e^{(-\lambda t)^{\beta}}$
- ☐ Failure rate:

$$\lambda(t) = \frac{f(t)}{R(t)} = \lambda \beta (\lambda t)^{\beta-1}$$

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□ No failures

The probability that within time interval [0, t+Δt] no failures occur is determined by multiplying the probability that until time t no failures have occurred (P₀(t)) and the probability that within [t, t+Δt] no failures occur (1-λ Δt):

$$P_0(t + \Delta t) = P_0(t)(1 - \lambda \Delta t) \Leftrightarrow \frac{P_0(t + \Delta t) - P_0(t)}{\Delta t} = -\lambda P_0(t)$$

For Δ t towards 0 one receives:

$$\lim_{\Delta t \to 0} \frac{P_0(t + \Delta t) - P_0(t)}{\Delta t} = \frac{d P_0(t)}{dt} = -\lambda P_0(t)$$

• $P_0(0) = 1$, since new systems (t=0) are always operational by definition. For a constant value of λ and $P_0(0) = 1$ the differential equation has the solution:

$$P_0(t) = e^{-\lambda t}$$

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The Poisson Distribution

The probability that a new system shows no failures until t is R(t)

$$R(t) = P_0(t) = e^{-\lambda t}$$

Using the definitions for F(t) and f(t), we get:

$$F(t) = 1 - R(t) = 1 - e^{-\lambda t}$$
 and $f(t) = \frac{dF(t)}{dt} = \lambda e^{-\lambda t}$

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□ Failures

The probability that within time interval [0, t+Δt] x failures occur can be determined as follows:

$$\begin{split} P_x\big(t+\Delta t\big) &= P_0\big(t\big) \, \left[P\big(x \;\; failures \;\; between \;\; t \;\; and \;\; t+\Delta t\big) \right] \\ &+\;\; \dots \\ &+\;\; P_{x-2}\big(t\big) \, \left[P\big(2 \;\; failures \;\; between \;\; t \;\; and \;\; t+\Delta t\big) \right] \\ &+\;\; P_{x-1}\big(t\big) \, \left[P\big(1 \;\; failure \;\; between \;\; t \;\; and \;\; t+\Delta t\big) \right] \\ &+\;\; P_x\big(t\big) \, \left[P\big(no \;\; failure \;\; between \;\; t \;\; and \;\; t+\Delta t\big) \right] \end{split}$$

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The Poisson Distribution

• Due to the precondition the probability to observe more than one failure in Δt is zero. Therefore we get:

$$\begin{split} P_x \left(t + \Delta t \right) &= P_{x-1} \left(t \right) \, \left[P \left(1 \right) \, failure \, \, between \, \, t \, \, \, and \, \, \, t + \Delta t \right) \right] \\ &+ P_x \left(t \right) \, \left[P \left(no \, \, \, failure \, \, \, between \, \, t \, \, \, and \, \, \, t + \Delta t \right) \right] \\ &= P_{x-1} \left(t \right) \left(\lambda \, \, \, \Delta t \right) + P_x \left(t \right) \left(1 - \lambda \, \, \, \, \Delta t \right) \\ &= P_x \left(t \right) - \left(\lambda \, \, \, \, \Delta t \right) \left[P_x \left(t \right) - P_{x-1} \left(t \right) \right] \\ &\Leftrightarrow \\ \frac{P_x \left(t + \Delta t \right) - P_x \left(t \right)}{\Delta t} &= -\lambda \left[P_x \left(t \right) - P_{x-1} \left(t \right) \right] \end{split}$$

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With ∆t approximating zero:

$$\lim_{\Delta t \to 0} \frac{P_x(t + \Delta t) - P_x(t)}{\Delta t} = \frac{dP_x(t)}{dt} = -\lambda [P_x(t) - P_{x-1}(t)]$$

The following term for P_x(t) is a solution for this differential equation

$$P_{x}(t) = \frac{(\lambda t)^{x} e^{-\lambda t}}{x!}$$

(Poisson Distribution)

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The Poisson Distribution

☐ This can be shown very easily

$$\frac{dP_{X}(t)}{dt} = \frac{d\frac{(\lambda t)^{x}e^{-\lambda t}}{x!}}{dt} = \frac{x(\lambda t)^{x-1}\lambda e^{-\lambda t} + (\lambda t)^{x}(-\lambda)e^{-\lambda t}}{x!}$$
$$= -\lambda \left[\frac{(\lambda t)^{x}e^{-\lambda t}}{x!} - \frac{(\lambda t)^{x-1}e^{-\lambda t}}{(x-1)!}\right] = -\lambda \left[P_{X}(t) - P_{X-1}(t)\right]$$

 $\ \square$ The probability $P_X(t)$ provides the correct value $P_0(t)$ also for the case that we treated separately before

$$\frac{dP_X(t)}{dt} = \frac{(\lambda t)^0 e^{-\lambda t}}{0!} = e^{-\lambda t} = P_0(t)$$

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□ $P_X(t)$ fulfills the boundary conditions for t = 0, i.e. $P_0(0) = 1$ and $P_X(0) = 0$, for $x \ge 1$. Furthermore the sum of the probabilities of all $x \ge 0$ for every $t \ge 0$ must be 1, i.e.

$$\sum_{x=0}^{\infty} P_x(t) = \sum_{x=0}^{\infty} \frac{(\lambda t)^x e^{-\lambda t}}{x!} = e^{-\lambda t} \sum_{x=0}^{\infty} \frac{(\lambda t)^x}{x!} = 1 \Rightarrow \sum_{x=0}^{\infty} \frac{(\lambda t)^x}{x!} \stackrel{!}{=} e^{-\lambda t}$$

 \square The specified sum on the left hand side of the equation is the power series of the exponential function on the right hand side. The Poisson Distribution thus fulfills the preconditions. If λ is constant, the mean value is $\mu(t)$ = λt . This is called a homogeneous Poisson Process. If λ is a function of time, the mean value is

$$\mu(t) = \int_{0}^{t} \lambda(\tau) d\tau \quad and \quad P_{x}(t) = \frac{\mu(t)^{x} e^{-\mu(t)}}{x!}$$

This is called a non-homogeneous Poisson Process (NHPP)

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Failure Times and Times between Failures

 \Box The time of failure i is T_i

 $\hfill\Box$ The time interval between failure (i - 1) and failure i is T_i

$$\Box T_i = \sum_{j=1}^{l} T_j^i, T_0 = 0$$

☐ M(t) is the number of failures at t

$$[M(t) \ge i] \Leftrightarrow |T_i| \le t$$

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Failure Times and Times between Failures

☐ The probability for j failures until time t is

$$P_{j}(t) = P[M(t) = j] = \frac{(\mu(t))^{j} e^{-\mu(t)}}{j!}$$

 $\hfill\Box$ The probability for at least i failures at t is

$$P[M(t) \ge i] = \sum_{j=i}^{\infty} \frac{(\mu(t))^j e^{-\mu(t)}}{j!} = P[T_i \le t]$$

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Musa's Execution Time Model

- □ A software system fails due to errors in the software randomly at t1,
 t2, ... (t here refers to execution time, i. e. CPU-seconds)
 - It is assumed that the number of failures observed in Δt is linearly proportional to the number of faults contained in the software at this time
 - $\mu(t)$ is the total number of failures for times $t \ge 0$
 - μ(t) is a limited function of t
 - The number of failures is a monotonic increasing function of t
 - At t=0 no failures have been observed yet: μ(0)=0
 - After very long execution time (t → ∞) the value µ(t) is equal to a. a is the total number of failures in infinite time. (There are also models where infinite numbers of failures are assumed to happen)

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Musa's Execution Time Model

- □ Model development
 - \blacksquare The number of failures observed in a time interval Δt is proportional to Δt and to the number of errors not yet detected

$$\mu(t + \Delta t) - \mu(t) = b[a - \mu(t)]\Delta t \Rightarrow \frac{\mu(t + \Delta t) - \mu(t)}{\Delta t} = ba - b\mu(t)$$

 $\hfill\Box$ With $\Delta t \to 0$ we get:

$$\frac{d\mu(t)}{dt} = ba - b\mu(t) = \mu'(t)$$

- □ With $\mu(0)$ =0 and $\mu(\infty)$ =a we get: $\mu(t) = a(1 e^{-bt})$
- □ The failure rate is: $\lambda(t) = \mu'(t) = abe^{-bt}$

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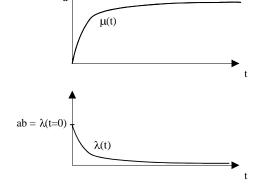


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 $\hfill \square$ The curve for the accumulated number of failures $\mu(t)$ approximates asymptotically the expected total number of failures a



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Musa's Elementary Execution Model

 \square The curve for the failure rate $\lambda(t)$ for t=0 starts at the initial failure rate $\lambda_0=$ ab and approximates asymptotically the value 0. The initial failure rate is proportional to the expected number of failures a, with the constant of proportionality b

$$\mu(t) = a\left(1 - e^{-bt}\right) = a\left(1 - e^{-\frac{\lambda_0}{at}}\right)$$

$$\lambda(t) = abe^{-bt} = \lambda_0 e^{-\frac{\lambda_0}{at}}$$

$$\mu(t) = a\left(1 - e^{-bt}\right) = a\left(1 - e^{-\frac{\lambda_0}{at}}\right) \quad and \quad \lambda(t) = abe^{-bt} = \lambda_0 e^{-\frac{\lambda_0}{at}} \implies e^{-bt} = \frac{\lambda(t)}{ab}$$

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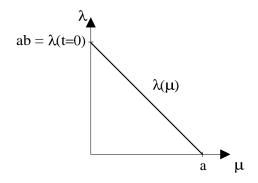


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Musa's Execution Time Model

$$\mu(t) = a \left(1 - \frac{\lambda(t)}{ab} \right) \Rightarrow \lambda(\mu) = b(a - \mu) = ab \left(1 - \frac{\mu}{a} \right) = \lambda_0 \left(1 - \frac{\mu}{a} \right)$$



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Musa's Execution Time Model

 $\hfill \square$ If λ is the present failure rate and a target λ_z is defined, $\Delta\mu$ additional failures will occur until this target is reached

$$\Delta \mu = \mu_z - \mu = a \left(1 - \frac{\lambda_z}{\lambda_0} \right) - a \left(1 - \frac{\lambda}{\lambda_0} \right) = a \left(\frac{\lambda - \lambda_z}{\lambda_0} \right)$$

 \square The additional time Δt until this target is reached is

$$\Delta t = t_z - t = -\frac{a}{\lambda_0} \left[\ln \left(\frac{\lambda_z}{\lambda_0} \right) - \ln \left(\frac{\lambda}{\lambda_0} \right) \right] = \frac{a}{\lambda_0} \left[\ln \left(\frac{\lambda}{\lambda_0} \right) - \ln \left(\frac{\lambda_z}{\lambda_0} \right) \right] = \frac{a}{\lambda_0} \ln \left(\frac{\lambda}{\lambda_z} \right)$$

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Musa's Execution Time Model

 \Box If

$$\mu(t) = a \left(1 - e^{-\frac{\lambda_0}{a}t} \right)$$

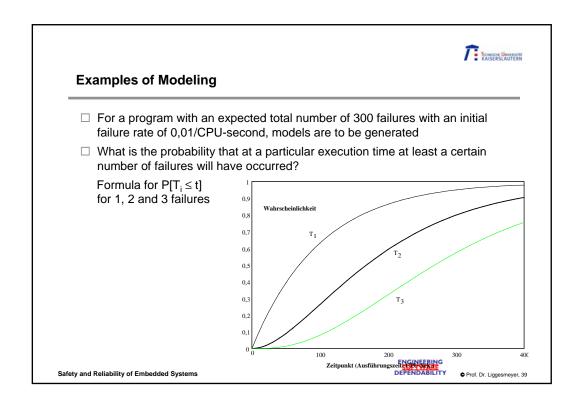
is inserted into the general equation of the Poisson distribution, we get:

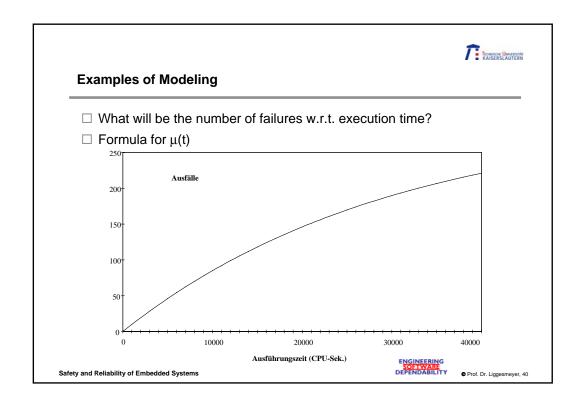
$$P[T_{i} \leq t] = \sum_{j=i}^{\infty} \frac{(\mu(t))^{j} e^{-\mu(t)}}{j!} = \sum_{j=i}^{\infty} \frac{\left[a\left(1 - e^{-\frac{\lambda_{0}}{a}t}\right)\right]^{j} e^{-a\left[1 - e^{-\frac{\lambda_{0}}{a}t}\right]}}{j!}$$

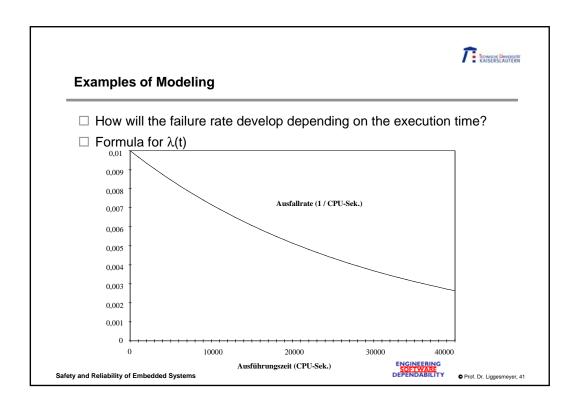
$$= e^{-a\left[1 - e^{-\frac{\lambda_{0}}{a}t}\right]} \sum_{j=i}^{\infty} \frac{\left[a\left(1 - e^{-\frac{\lambda_{0}}{a}t}\right)\right]^{j}}{j!}$$

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Determination of Model Parameters

□ Least squares

Target: Define parameters in such a way that the sum of the squares of the deviations between the calculated and the observed values becomes minimal. If F_i refers to the value of the empirical distribution function at point t_i, the following term is to be minimized:

$$\Delta = \sum_{i=1}^{n} \Delta_{i}^{2} = \sum_{i=1}^{n} (F(t_{i}) - F_{i})^{2}$$

☐ Maximum-Likelihood-Method

 Target: Choose parameters in such a way that the probability is maximized to produce a "similar" observation to the present observation.
 The probability density has to be known

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Determination of Model Parameters Least squares

☐ Target: Define parameters in such a way that the sum of the squares of the deviations between the calculated and the observed values becomes minimal. If F_i refers to the value of the empirical distribution function at point t_i, the following term is to be minimized:

$$\Delta = \sum_{i=1}^{n} \Delta_{i}^{2} = \sum_{i=1}^{n} (F(t_{i}) - F_{i})^{2}$$

 \square For the exponential distribution we get:

$$\Delta_{\text{exp}} = \sum_{i=1}^{n} \Delta_{i_{\text{exp}}}^{2} = \sum_{i=1}^{n} (F(t_{i}) - F_{i})^{2} = \sum_{i=1}^{n} (1 - e^{-\lambda t_{i}} - F_{i})^{2}$$

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Determination of Model Parameters Least squares



 $\hfill\Box$ The value λ that minimizes this term is to be determined

$$\frac{d\Delta_{\exp}}{d\lambda} = \frac{d\left(\sum_{i=1}^{n} \Delta_{i_{\exp}}^{2}\right)}{d\lambda} = \sum_{i=1}^{n} 2\left(1 - e^{-\lambda t_{i}} - F_{i}\right) t_{i} e^{-\lambda t_{i}}$$

 \square The value $\hat{\lambda}$ is calculated by determining the root

$$\sum 2(1 - e^{-\hat{\lambda}t_i} - F_i)t_i e^{-\hat{\lambda}t_i} = 0$$

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Determination of Model Parameters Least squares

☐ Sometimes numerical method must be used for this task. A Newtonian iteration provides the following results for the Exponential Distribution

$$\lambda_{n+1} = \lambda_n - \frac{f(\lambda_n)}{\frac{df(\lambda_n)}{d\lambda}}$$

$$\Box \text{ with: } f_{\exp}(\lambda) = \sum_{i=1}^{n} 2(1 - e^{-\lambda t_i} - F_i) t_i e^{-\lambda t_i}$$

 \Box For the failure times of the table on page 10 the search for zero points according to the Newtonian iteration provides a value $\hat{\lambda} \approx 3,9326702$ * $10^{-5}/h$ for the exponential distribution

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Determination of Model Parameters Maximum-Likelihood-Method



- ☐ Target: Choose parameters in such a way that the probability is maximized to produce a "similar" observation to the present observation
- ☐ Precondition: Probability density has to be known
- □ Likelihood function
 - Product of the densities at the observed failure times
 - The value is proportional to the probability to observe failure times that do not exceed the deviation Δt w.r.t. the present observation
 - It is a function of the distribution function's parameters that are to be determined
 - Example:

The parameter λ of the exponential distribution is to be determined with the Maximum-Likelihood-Method

$$F(t)=1-e^{-\lambda t}, f(t)=\lambda e^{-\lambda t}$$

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Determination of Model Parameters Maximum-Likelihood-Method

With n observed failure times $t_1, ..., t_n$ we get the Likelihood Function:

$$\begin{split} L(\lambda, t_1, \dots, t_n) &= f(\lambda, t_1) f(\lambda, t_2) \dots f(\lambda, t_n) = \lambda e^{-\lambda t_1} \lambda e^{-\lambda t_2} \dots \lambda e^{-\lambda t_n} \\ &= \lambda^n e^{-\lambda (t_1 + t_2 + \dots + t_n)} = \lambda^n e^{-\lambda \sum_{i=1}^n t_i} \end{split}$$

Due to the monotonicity of the logarithmic function, \boldsymbol{L} und $\ln \boldsymbol{L}$ have identical maxima

$$\ln L(\lambda, t_1, \dots, t_n) = n \ln \lambda - \lambda \sum_{i=1}^n t_i$$

In order to calculate the value $\hat{\lambda}$ that maximizes the Likelihood Function, the derivation according to λ must be determined

$$\frac{d(\ln L(\lambda,t_1,\ldots,t_n))}{d\lambda} = \frac{n}{\lambda} - \sum_{i=1}^n t_i$$

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Determination of Model Parameters Maximum-Likelihood-Method



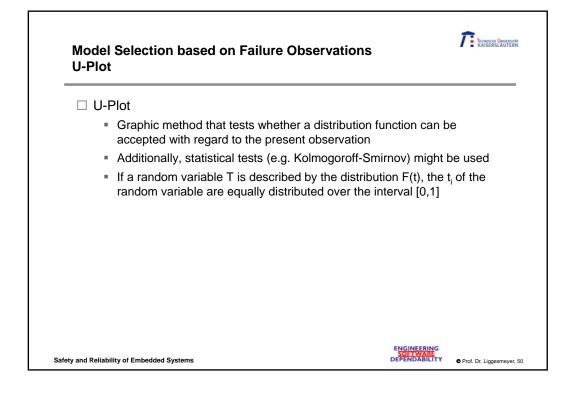
 \Box $\hat{\lambda}$ is the root. For the exponential distribution we get:

$$\frac{n}{\lambda} - \sum_{i=1}^{n} t_i \stackrel{!}{=} 0 \iff \hat{\lambda} = \frac{n}{\sum_{i=1}^{n} t_i}$$

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Model Selection based on Failure Observa	tions	
☐ U-Plot-Method		
☐ Prequential-Likelihood-Method		
☐ Holdout-Evaluation		
	ENGINEERING SOFTWARE	
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Model Selection based on Failure Observations U-Plot

- The n values U_i are charted in a U-Plot as follows
 - The values U_i are used as y-values in such a way that the value U_i with the position j is attributed to the x-value $^{ij}\!\!/_n$
 - If the values U_i are approximately equally distributed, the applied points are located "near by" the function y=x, for $0 \le x \le 1$

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Model Selection based on Failure Observations U-Plot



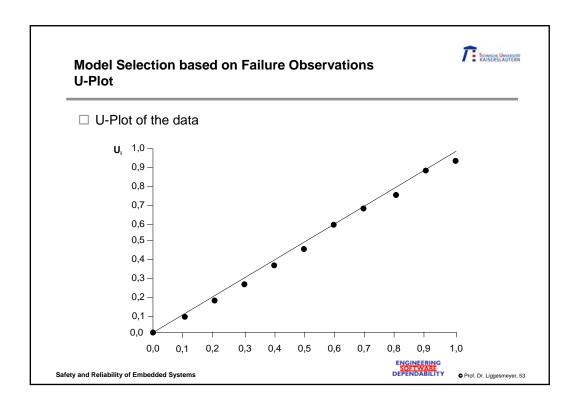
□ Example

T _i (h)	2810	5411	8701	13130	17327	24899	31230	40006	59880	80017
n/ _N	0,1	0,2	0,3	0,4	0,5	0,6	0,7	0,8	0,9	1,0
F(t)	0,094	0,174	0,264	0,371	0,457	0,585	0,668	0,765	0,879	0,941

The values presented in the table for F(t) are the $U_{\rm i}$ according to the definition stated above

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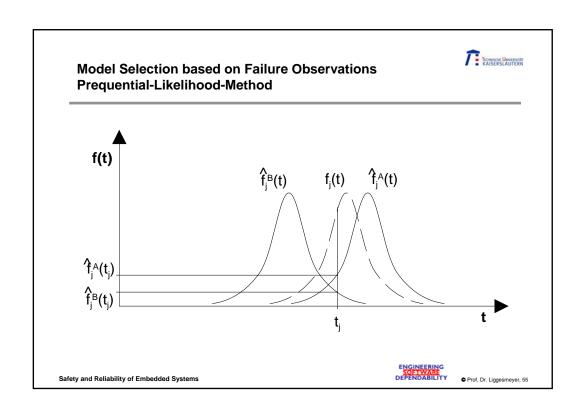
Model Selection based on Failure Observations Prequential-Likelihood-Method



- ☐ The Prequential-Likelihood-Method compares the suitability of two distribution functions under consideration with regard to a given failure observation
- ☐ It is based on the following approach
 - The failure interval t_j is a realization of a random variable with the distribution F_j(t) and the density f_j(t)
 - F_i(t) and f_i(t) are unknown
 - The densities of the distribution functions A and B ($\hat{f}_j^A(t)$ resp. $\hat{f}_j^B(t)$) can be determined based on the failure intervals $\mathbf{t_1}$, ..., $\mathbf{t_{j-1}}$
 - If the distribution A is more suitable than the distribution B, it can be expected that the value $\hat{f}_j^A(t_j)$ is greater than the value $\hat{f}_j^B(t_j)$
 - The quotient $\frac{\hat{f}_{j}^{A}(t_{j})}{\hat{f}_{i}^{B}(t_{j})}$ will be greater than 1

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Model Selection based on Failure Observations Prequential-Likelihood-Method



 $\hfill \square$ If this analysis is done for every observed failure time interval t_j we get the so-called Prequential-Likelihood-Ratio concerning the distributions A and B

$$PLR_{i}^{AB} = \prod_{j=s}^{j=i} \frac{\hat{f}_{j}^{A}(t_{j})}{\hat{f}_{j}^{B}(t_{j})}$$

☐ If A is more appropriate than B with regard to the present failure data, the PLR shows a rising tendency

□ Example

 We compare the exponential distribution and the normal distribution based on the data from the table on page 10 using the Prequential-Likelihood-Method

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Model Selection based on Failure Observations Prequential-Likelihood-Method

The parameters of the distributions are determined using a Maximum-Likelihood-Approach. For the exponential distribution, we get:

$$\hat{\lambda}_j = \frac{j-1}{\sum\limits_{k=1}^{j-1} t_k}$$

- For the normal distribution we get: $f(\tau, \sigma^2, t) = \frac{1}{\sqrt{(2\pi\sigma^2)}} e^{-(t-\tau)^2/2\sigma^2}$
- The parameters according to the Maximum-Likelihood-Method are:

$$\hat{\tau}_{j} = \frac{\sum_{k=1}^{j-1} t_{k}}{j-1}; \hat{\sigma}_{j}^{2} = \frac{1}{j-1} \sum_{k=1}^{j-1} (t_{k} - \tau)^{2}$$

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Model Selection based on Failure Observations Prequential-Likelihood-Method

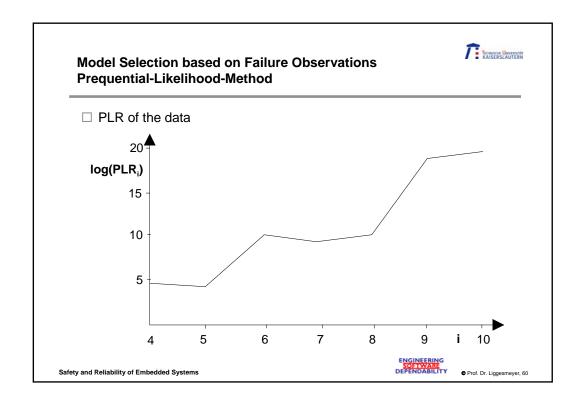


- The following table shows the densities of the exponential distribution and the normal distribution for the arrival time intervals t_i based on the failure times T₁ to T_{i-1} from the table on page 10
- The calculation starts with i = 4. In addition the logarithm of the quotient of the densities and the logarithm of the PLR_i is contained in the table
- The rising of the PLR underlines that the assumption of exponentially distributed arrival times for the present data makes more sense than the assumption of normally distributed arrival times

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Model Selection based on Failure Observations Prequential-Likelihood-Method										Tronson University KAISERSLAUTERN	
i	1	2	3	4	5	6	7	8	9	10	
T _i (h)	2810	5411	8701	13130	17327	24899	31230	40006	59880	80017	
t _i (h)	2810	2601	3290	4429	4197	7572	6331	8776	19874	20137	
f _i Exp/ 10 ⁻⁶				74,9	84,8	32,5	52,4	31,3	3,8	7,3	
f _i Norm/ 10 ⁻⁹				1,1	244558	0,076	101787	10096	0,000008	2362	
log (f _i Exp/ f _i Norm)				4,83	-0,46	5,63	-0,29	0,49	8,67	0,49	
log (PLR _i)				4,83	4,37	10,00	9,71	10,20	18,87	19,36	





Model Selection based on Failure Observations Holdout Evaluation

□ Approach

- Only parts of the failure data are used for model calibration. The remaining data are used to judge the prediction quality of the calibrated model
- If an exponential distribution and a Weibull distribution are calibrated to the first 6 failure times (table p. 10) using a Least-Squares-Algorithm, we get the following results:
- \blacksquare Exponential distribution: $F(t)_{\rm exp} = 1 e^{-3.89292*10^{-5}t_i/h}$
- $\text{Weibull distribution:} \qquad F(t)_{\rm exp} = 1 e^{-\frac{1}{1.552005*10^4}t_i^{0.94750}/h}$

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Model Selection based on Failure Observations Holdout Evaluation



- \square The Weibull distribution has as expected a better adjustment to the failure times T_1 to T_6 . The sum of the deviation squares for the first 6 failure times is 0,000459 compared to 0,000790 in the exponential distribution
- ☐ The prediction quality of the Weibull distribution is however worse than that of the exponential distribution. The sum of the deviation squares for the failure times T₇ to T₁₀ is 0,00446 for the Weibull distribution; for the exponential distribution is only 0,00210. We might prefer to use the exponential distribution in order to avoid "over-calibration"

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